

## SOLUTIONS—ASSIGNMENT 9

Chapter 6: P 1. The sample space is  $S = \{(m, n) : m, n \text{ integers } 1, 2, \dots, 6\}$ , so  $|S| = 36$ .

(a) Here  $X(m, n) = \max\{m, n\}$  and  $Y(m, n) = m + n$ . The joint range of  $(X, Y)$  is all pairs of integers  $(x, y)$  with  $x + 1 \leq y \leq 2x$  and  $1 \leq x \leq 6$ ,  $2 \leq y \leq 12$ . If  $y = 2x$ , then only one point  $(x, x) \in S$  gives this value, so  $p(x, 2x) = 1/36$ . If  $x < y < 2x$  then the points  $(x, y - x)$  and  $(y - x, x)$  in  $S$  both give this value for  $(X, Y)$ . So  $p(x, y) = 2/36$  in this case. Here is a table:

$X \setminus Y$	2	3	4	5	6	7	8	9	10	11	12
1	1/36										
2		1/18	1/36								
3			1/18	1/18	1/36						
4				1/18	1/18	1/18	1/36				
5					1/18	1/18	1/18	1/18	1/36		
6						1/18	1/18	1/18	1/18	1/18	1/36

(b) Now  $X(m, n) = m$  and  $Y(m, n) = \max\{m, n\}$ . The joint range of  $(X, Y)$  is all pairs of integers  $(x, y)$  with  $x \leq y$  and  $x, y = 1, \dots, 6$ . For  $x = y$ , the points  $(x, 1), (x, 2), \dots, (x, x)$  of  $S$  give this value for  $(X, Y)$ , so  $p(x, x) = x/36$ . For  $x < y$ , the only point of  $S$  giving this value is  $(x, y)$ . So  $p(x, y) = 1/36$ . See the table below.

(c) Now  $X(m, n) = \min\{m, n\}$  and  $Y(m, n) = \max\{m, n\}$ . The joint range of  $(X, Y)$  is the same as in (b). If  $x = y$ , then the only point of  $S$  giving this value for  $(X, Y)$  is  $(x, y)$ . So  $p(x, y) = 1/36$ . If  $x < y$  then the points  $(x, y)$  and  $(y, x)$  of  $S$  both give this value of  $(X, Y)$ . Here is the table:

$Y \setminus X$	1	2	3	4	5	6
1	1/36					
2	1/36	2/36				
3	1/36	1/36	3/36			
4	1/36	1/36	1/36	4/36		
5	1/36	1/36	1/36	1/36	5/36	
6	1/36	1/36	1/36	1/36	1/36	6/36

Table for 1(b)

$Y \setminus X$	1	2	3	4	5	6
1	1/36					
2	2/36	1/36				
3	2/36	2/36	1/36			
4	2/36	2/36	2/36	1/36		
5	2/36	2/36	2/36	2/36	1/36	
6	2/36	2/36	2/36	2/36	2/36	1/36

Table for 1(c)

6. Think of testing all 5 transistors in order. There are  $\binom{5}{2}$  ways to choose the two positions in the testing order for the defective transistors, and all of these choices are equally likely; each choice corresponds to a possible value of  $N_1$  and  $N_2$ , with  $N_1$  the position of the first defective in the testing order and  $N_1 + N_2$  the position of the second. These numbers must satisfy  $1 \leq N_1 < N_1 + N_2 \leq 5$  or  $1 \leq N_1 \leq 4$ ,  $1 \leq N_2 \leq 5 - N_1$ . All  $\binom{5}{2}$  pairs  $(N_1, N_2)$  satisfying these inequalities are equally likely; each has probability  $1/\binom{5}{2} = 1/10$ .

8. (a) As usual, we use the principle that the total probability is 1:

$$\begin{aligned} 1 &= \iint f(x, y) dx dy = c \int_0^\infty \int_{-y}^y (y^2 - x^2)e^{-y} dx dy \\ &= c \int_0^\infty e^{-y} [y^2 x - x^3/3]_{-y}^y dy = \frac{4c}{3} \int_0^\infty y^3 e^{-y} dy = 8c. \end{aligned}$$

So  $c = 1/8$ .

(b) Since  $Y > 0$ ,  $f_Y(y) = 0$  for  $y < 0$ . For  $y \geq 0$ ,

$$f_Y(y) = \int_{-\infty}^\infty f(x, y) dx = \frac{1}{8} \int_{-y}^y (y^2 - x^2)e^{-y} dx = \frac{1}{6} y^3 e^{-y}.$$

For any  $x$ ,  $f(x, y) = 0$  unless  $y > |x|$ . So

$$f_X(x) = \frac{1}{8} \int_{|x|}^{\infty} (y^2 - x^2)e^{-y} dy = \frac{1}{8} \left[ (-y^2 - 2y - 2 + x^2)e^{-y} \right]_{|x|}^{\infty} = \frac{1}{4} (|x| + 1)e^{-|x|}.$$

(c) Since  $f_X$  is even ( $f_X(x) = f_X(-x)$ ),  $E[X] = 0$  by symmetry:

$$\begin{aligned} E[X] &= \int_{-\infty}^0 xf(x) dx + \int_0^{\infty} xf(x) dx = \int_{\infty}^0 (-y)f(-y) (-dy) + \int_0^{\infty} xf(x) dx \\ &= - \int_0^{\infty} yf(y) dy + \int_0^{\infty} xf(x) dx. \end{aligned}$$

where we have made the substitution  $y = -x$ .

9. (a)  $f(x, y)$  is clearly non-negative. The other requirement for a density is that it be normalized, which we can check:

$$\frac{6}{7} \int_0^1 \int_0^2 \left( x^2 + \frac{xy}{2} \right) dy dx = \frac{6}{7} \int_0^1 (2x^2 + x) dx = \frac{6}{7} \left( \frac{2}{3} + \frac{1}{2} \right) = 1.$$

(b) For  $0 \leq x \leq 1$ ,  $f_X(x) = \frac{6}{7} \int_0^2 \left( x^2 + \frac{xy}{2} \right) dy = \frac{6}{7} (2x^2 + x)$ .

(c)

$$P\{X > Y\} = \frac{6}{7} \int_0^1 \int_0^x \left( x^2 + \frac{xy}{2} \right) dy dx = \frac{6}{7} \int_0^1 \frac{5x^3}{4} dx = \frac{15}{56}.$$

(d)

$$P\{X < \frac{1}{2}\} = \int_0^{1/2} f_X(x) dx = \frac{6}{7} \left[ \frac{2x^3}{3} + \frac{x^2}{2} \right]_0^{1/2} = \frac{5}{28}.$$

So

$$\begin{aligned} P\left(\left\{Y > \frac{1}{2}\right\} \mid \left\{X < \frac{1}{2}\right\}\right) &= \frac{28}{5} P\left(\left\{Y > \frac{1}{2}\right\} \text{ and } \left\{X < \frac{1}{2}\right\}\right) \\ &= \frac{28}{5} \frac{6}{7} \int_0^{1/2} \int_{1/2}^2 \left( x^2 + \frac{xy}{2} \right) dy dx \\ &= \frac{24}{5} \int_0^{1/2} \left( \frac{3x^2}{2} + \frac{15x}{16} \right) dx \\ &= \frac{24}{5} \left( \frac{1}{16} + \frac{15}{128} \right) = \frac{69}{80}. \end{aligned}$$

(d)

$$E[X] = \int_{-\infty}^{\infty} xf_X(x) dx = \frac{6}{7} \int_0^1 x(2x^2 + x) dx = \frac{5}{7}.$$

(e)

$$f_Y(y) = \frac{6}{7} \int_0^1 \left( x^2 + \frac{xy}{2} \right) dx = \frac{2}{7} + \frac{3y}{14}; \quad E[Y] = \int_0^2 y \left( \frac{2}{7} + \frac{3y}{14} \right) dy = \frac{8}{7}$$

10. (a)  $P\{X < Y\} = \int_0^{\infty} \int_0^y e^{-(x+y)} dx dy = \int_0^{\infty} e^{-y} [1 - e^{-y}] dy = 1 - 1/2 = 1/2$  (this could be seen without calculation by symmetry in  $X$  and  $Y$ ).

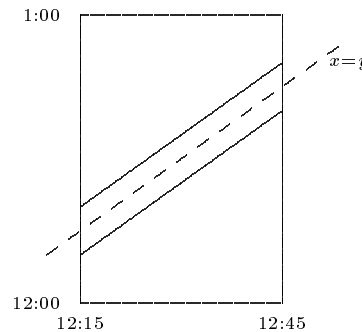
(b) Since  $f(x, y) = e^{-x} \cdot e^{-y}$ ,  $X$  and  $Y$  are independent and each has exponential distribution with parameter  $\lambda = 1$ . So  $P\{X < a\} = 1 - e^{-a}$ .

13. If  $X$  and  $Y$  are the arrival times of the man and woman, respectively, in minutes after noon, then  $(X, Y)$  is uniformly distributed on the rectangle  $15 \leq x \leq 45$ ,  $0 \leq y \leq 60$ . We can thus determine probabilities geometrically, as ratios of areas.

(a) The set in which the early arriver must wait at most 5 minutes,  $\{|x - y| \leq 5\}$ , is a parallelogram of altitude 30 and base 10 (shown in the figure), so

$$P\{|X - Y| \leq 5\} = (10 \times 30)/(30 \times 60) = 1/6.$$

(b) It seems clear from the symmetry between the man and the woman that we should have  $P\{X < Y\} = 1/2$ , and this is easily verified by computing the area of the trapezoid lying below the dashed line in the figure.



15. (a) Since  $\int \int_R 1 \, dx \, dy = (\text{area of } R)$ , we have  $\int \int_R \frac{1}{(\text{area of } R)} \, dx \, dy = 1$ .

(b) Here  $R$  has area 4 so the joint density is  $f(x, y) = 1/4$  if  $-1 \leq x, y \leq 1$ ,  $f(x, y) = 0$  otherwise. Thus

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) \, dy = \begin{cases} (1/4) \int_{-1}^1 dy = 1/2, & \text{if } -1 \leq x \leq 1, \\ 0, & \text{otherwise.} \end{cases}$$

Clearly  $X$  is uniform on  $(-1, 1)$ . Similarly,  $Y$  is uniform on  $(-1, 1)$ . Moreover,  $X$  and  $Y$  are independent because  $f(x, y) = f_X(x)f_Y(y)$ .

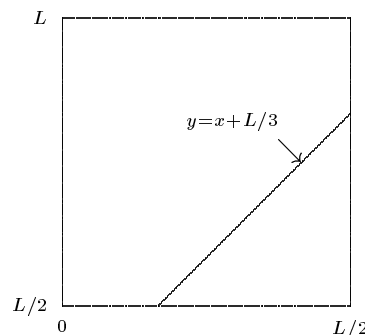
(c) Let  $A$  be the region inside the circle of radius 1, centered at the origin. Then

$$P\{X^2 + Y^2 \leq 1\} = P\{(X, Y) \in A\} = \frac{|A|}{4} = \pi/4.$$

18.  $X$  and  $Y$  are uniformly distributed on the square  $0 \leq X \leq L/2$ ,  $L/2 \leq Y \leq L$ , so we compute probabilities as ratios of areas. The set where  $Y - X > L/3$  is the complement of a triangle with base and height  $L/3$ , so

$$P\{Y > X + L/3\} = 1 - \frac{(L/3)^2(1/2)}{(L/2)^2} = \frac{7}{9}.$$

See the Figure.



20. (a) Integrating with respect to  $y$  yields  $f_X(x) = xe^{-x}$  for  $x \geq 0$ , and similarly  $f_Y(y) = e^{-y}$  for  $y \geq 0$ ;  $f_X$  and  $f_Y$  are zero if their arguments are negative. Since  $f_{X,Y}(x, y) = f_X(x) \cdot f_Y(y)$ ,  $X$  and  $Y$  are independent.

(b) Here

$$f_X(x) = \begin{cases} 2(1-x) & 0 < x < 1, \\ 0 & \text{otherwise;} \end{cases}$$

$$f_Y(y) = \begin{cases} 2y & 0 < y < 1, \\ 0 & \text{otherwise,} \end{cases}$$

so  $X$  and  $Y$  are not independent.

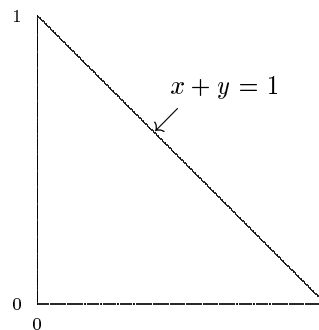
21. (a) Clearly  $f(x, y) \geq 0$ ; we check normalization:

$$\int_0^1 \int_0^{1-x} 24xy \, dy \, dx = 12 \int_0^1 x(1-x)^2 \, dx = 1.$$

(b)  $f_X(x) = \int_0^{1-x} 24xy \, dy = 12x(1-x)^2$ ,  $0 \leq x \leq 1$ ;

$$E[X] = \int_0^1 xf(x) \, dx = 2/5.$$

(c) Just as in (b) (or by symmetry):  $E[Y] = 2/5$ .



25. For each individual, the probability that he or she arrives in the first hour is  $10^{-6}$ . Thus the number  $N$  of people that arrive in the first hour is binomial with parameters  $n = 10^6$ ,  $p = 10^{-6}$ . This is well approximated by a Poisson distribution with  $\lambda = np = 1$ , so  $P(N = i) \approx e^{-1}/i!$ .

27. By independence,  $X$  and  $Y$  have joint density

$$f(x, y) = f_X(x)f_Y(y) = \begin{cases} e^{-y} & \text{if } 0 \leq x \leq 1, y \geq 0 \\ 0 & \text{otherwise.} \end{cases}$$

(a) If  $Z = X + Y$ , then

$$\begin{aligned} F_Z(z) &= P\{X + Y \leq z\} = P\{Y \leq z - X\} \\ &= \int \int_{y \leq z-x} f(x, y) dy dx. \end{aligned}$$

Setting up the limits on this integral depends on whether  $z > 1$  or  $z < 1$ . If  $z < 1$ ,

$$F_Z(z) = \int_0^z \int_0^{z-x} e^{-y} dy dx = \int_0^z (1 - e^{x-z}) dx = z - 1 + e^{-z},$$

while if  $z > 1$ ,

$$F_Z(z) = \int_0^1 \int_0^{z-x} e^{-y} dy dx = \int_0^1 (1 - e^{x-z}) dx = 1 - e^{1-z} + e^{-z}.$$

Thus

$$f_Z(z) = \begin{cases} 1 - e^{-z}, & \text{for } 0 < z < 1; \\ e^{-(z-1)} - e^{-z} & \text{for } z \geq 1. \end{cases}$$

(b) If  $Z = X/Y$ , then

$$\begin{aligned} F_Z(z) &= P\{Y > X/z\} = \int_0^1 \int_{x/z}^{\infty} e^{-y} dy dx \\ &= \int_0^1 (e^{-x/z}) dx = z[1 - e^{-1/z}]. \end{aligned}$$

So

$$\begin{aligned} f_Z(z) &= \frac{d}{dz} P(Z \leq z) \\ &= \begin{cases} (1 - e^{-1/z}) + \frac{e^{-1/z}}{z}, & \text{for } 0 < z, \\ 0, & \text{otherwise.} \end{cases} \end{aligned}$$

